Influence of Location, Population, and Climate on Building Damage and Fatalities due to Australian Bushfire: 1925–2009

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ABSTRACT

This study reevaluates the history of building damage and loss of life due to bushfire (wildfire) in Australia since 1925 in light of the 2009 Black Saturday fires in Victoria in which 173 people lost their lives and 2298 homes were destroyed along with many other structures. Historical records are normalized to estimate building damage and fatalities had events occurred under the societal conditions of 2008/09. There are relationships between normalized building damage and the El Niño–Southern Oscillation and Indian Ocean dipole phenomena, but there is no discernable evidence that the normalized data are being influenced by climatic change due to the emission of greenhouse gases. The 2009 Black Saturday fires rank second in terms of normalized fatalities and fourth in terms of normalized building damage. The public safety concern is that, of the 10 years with the highest normalized building damage, the 2008/09 bushfire season ranks third, behind the 1925/26 and 1938/39 seasons, in terms of the ratio of normalized fatalities to building damage. A feature of the building damage in the 2009 Black Saturday fires in some of the most affected towns—Marysville and Kinglake—is the large proportion of buildings destroyed either within bushland or at very small distances from it (<10 m). Land use planning policies in bushfire-prone parts of this country that allow such development increase the risk that bushfires pose to the public and the built environment.

1. Introduction

Widely heralded in the media as Australia’s worst natural disaster (e.g., Rule 2009), the 7 February 2009 Black Saturday bushfires (wildfires) in Victoria were but the most recent reminder of the potential for natural hazards to impact Australian communities (Crompton and McAneney 2008). Fueled by record high temperatures and high winds in the midst of a protracted drought, the Black Saturday fires claimed 173 lives and 2298 houses (Victorian Bushfires Royal Commission 2009) as well as numerous other structures, including schools and police stations. This paper attempts to place these most recent bushfire impacts into a historical context.

Following a method analogous to Crompton and McAneney (2008) and other recent work (Bouwer 2010), this paper asks: What would have been the impact of past bushfires if they were to recur under current societal conditions? Without accounting for the known influence societal factors have on disaster records, it is impossible to know whether the devastation inflicted by the Black Saturday fires was truly anomalous, whether this provides a glimpse of the future under expected changes in climate, and what policy changes might prove effective in reducing the impact of future disasters. In examining such questions, we shall also reevaluate work undertaken before the Black Saturday fires (e.g., McAneney et al. 2009) and present some patterns of building destruction in these particular bushfires.

Despite claims that the Black Saturday fires were Australia’s greatest natural disaster (e.g., Rule 2009), several previous events have been more destructive in
terms of loss of life and property damage, even before the societal influence has been accounted for: in 1974, Cyclone Tracy laid waste the city of Darwin, demolishing about 3700 dwellings and damaging another 3300 to the point that only 6% of the building stock was left habitable (Walker 1975); in 1899, Cyclone Mahina, a category-5 tropical cyclone, claimed about 410 lives; and the heat wave that preceded the 1939 Black Friday bushfires in Victoria is blamed for 438 deaths (from the research organization Risk Frontiers’s “PerilAUS” natural disaster database, described below in section 2a; Blong 2004; McAneney et al. 2009; Haynes et al. 2010). Of the more extreme bushfires, 1694 houses were lost in the 1983 Ash Wednesday fires in Victoria and South Australia (PerilAUS; Blong 2004; McAneney et al. 2009; Haynes et al. 2010), and, although we have been unable to verify this independently, were reportedly destroyed in the 1898 Red Tuesday fires in Victoria (State Government of Victoria 2003, p. 10). Regardless of its ranking in terms of numbers of fatalities and property damage, the extreme impacts in the Black Saturday fires warrant critical examination. This same sentiment led the Victorian state government to form a royal commission with wide executive powers to scrutinize all aspects of bushfire management leading up to and during the bushfires (Victorian Bushfires Royal Commission 2009).

The process of adjusting time series of disaster losses for changes in population, wealth and inflation and, in some cases, improved construction standards is known as normalization and has been applied in a wide range of locales for a range of phenomena (e.g., Pielke and Landsea 1998; Pielke et al. 2008; Crompton and McAneney 2008; Zhang et al. 2009; Barredo 2009, 2010; Vranes and Pielke 2009). Accounting for inflation/deflation is necessary because the value of a currency changes over time while increases in population and wealth mean more people and property are located in exposed areas.

In respect to Australian bushfire, McAneney et al. (2009) argued that the stability over the last century of exceedance loss statistics for building damage suggested that it was premature to conclude that a signal of greenhouse-gas emissions was present. The authors contend that, given that these loss statistics had proved so stable in the face of the vast societal changes that took place over the twentieth century, any greenhouse-gas signal cannot be large or significant. This study revisits this question using a different approach by explicitly accounting for these societal changes.

Whereas a greenhouse gas–driven climatic-change signal has thus far not been detected in normalized disaster loss records for a wide range of perils in locations around the world [see review by Bouwer (2010) and references therein], and is unlikely to be detected in at least storm and flood losses in the near future (Höppe and Pielke 2006), patterns of behavior characteristic of meteorological cycles such as El Niño–Southern Oscillation (ENSO) have been identified in normalized Atlantic Ocean hurricane damages (Pielke and Landsea 1999). ENSO and another coupled ocean–atmosphere oscillation, the Indian Ocean dipole (IOD), are also known to influence the weather and climate of eastern Australia (McBride and Nicholls 1983; Power et al. 2006; Ashok et al. 2003; Cai et al. 2009b); the former oscillation is in the equatorial Pacific Ocean and the latter in the Indian Ocean.

An El Niño (La Niña) phase of the ENSO cycle refers to the situation in which sea surface temperatures in the central to eastern Pacific Ocean are significantly warmer (cooler) than the long-term average, whereas a positive IOD (pIOD) event is when the eastern Indian Ocean is cooler than normal and the western Indian Ocean is anomalously warmer (Saji et al. 1999). El Niño events increase the chance of drought along eastern Australia (Kiern and Franks 2004) and bushfire (Williams and Karoly 1999), whereas La Niña events often presage widespread increases in rainfall (Power et al. 2006) and chance of flooding (Kiern et al. 2003). Ummenhofer et al. (2009) showed that a lack of negative IOD (nIOD) events was strongly related to drought in southeastern Australia, and Cai et al. (2009a) report a link between pIOD events and enhanced bushfire risk over Victoria. Moreover, Cai et al. (2009a) found that pIOD events were more effective than El Niño events in preconditioning Victorian bushfires, a robust result that was not conditional on the definitions adopted for each. This paper will examine the relationships between ENSO and the IOD and normalized bushfire building damage in Australia.

The remainder of this paper is structured as follows: we begin with a description of Risk Frontiers’ PerilAUS inventory of Australian bushfire building damage and the bushfire fatality database of Haynes et al. (2010). The normalization methods, ENSO and IOD definitions, and the method used to examine patterns of building damage in the Black Saturday fires for two of the most severely impacted towns, Marysville and Kinglake, are then detailed. We then present key results, including those from two historic case studies (the 1967 Hobart fires and the 1983 Ash Wednesday fires) used to “ground truth” the normalization method. The paper concludes with a discussion of results and some implications for public policy with regard to bushfire in Australia.

2. Data and methods

a. Bushfire building damage and fatality data

The current study draws upon Risk Frontiers’ databases of natural disasters in Australia (hereinafter referred to as
b. Normalizing house equivalents

To normalize bushfire building damage (HE) records to current societal conditions we simply convert the HE in bushfire year $i$ ($HE_i$) to bushfire year 2008 numbers ($HE_{08}$) as follows:

\[
HE_{08} = HE_i \times N_{i,j}
\]

where $N_{i,j}$ is the dwelling number factor defined as the ratio of the number of dwellings in bushfire year 2008 in state or territory $j$ to those present in bushfire year $i$. The number of dwellings in each state or territory is reported in the census of population and housing and/or year books [available from the Australian Bureau of Statistics (ABS) online at http://www.abs.gov.au]. A dwelling is defined as a structure intended for human habitation—normally a house, flat, caravan, and so on—but also includes hotels, prisons, hospitals, and so on that were occupied on census night. National censuses were undertaken irregularly until 1961 and at 5-yearly intervals since. Linear interpolation was used to determine the number of dwellings for years between census years, and the 2007 and 2008 bushfire year numbers were estimated by extrapolating from the 2001 and 2006 figures. Growth in the number of dwellings is assumed as a proxy for growth in HE.

Equation (1) ignores any explicit correction for inflation and wealth as measured in economic terms. The HE representation avoids the need for an inflation adjustment; whether an adjustment for increasing economic wealth is required is less obvious. An argument for its inclusion stems from the manifest increase in the average size of Australian dwellings over time: for example, the average increase in the average number of bedrooms per dwelling between 1976 and 2006 was 0.3% per year (from the ABS data). If this rate of increase had held constant over the entire analysis period, then the average dwelling size would have increased by 28% between 1925 and 2008. On the other hand, we expect most of that increase has been implicitly accounted for in the manner by which the HE data were derived: if, by way of example, we imagine a hypothetical bushfire event in which 100 houses were destroyed, then we assume that this equates to 100 HE whether the event occurred in 1930 or 1990. Although Blong (2003) differentiates between small, median, and large houses based on floor area, this level of detail is not often included in the source documents and so, for most building types, numbers of HE were based on a single (median) size of each building type. This being the case, we have chosen not to further adjust the HE data for changes in wealth; however, any adjustment of economic losses would also require both an inflation and economic wealth adjustment.

c. Normalizing fatalities

Bushfire-related fatalities $F$ are normalized in a similar manner to HE under the assumption that fatalities change in proportion to population (Pielke et al. 2003; Vranes and Pielke 2009):
\[ F_{08} = F_i \times P_{i,j} \]

where \( P_{i,j} \) is the population factor defined as the ratio of the population in bushfire year 2008 in state or territory \( i \) to the population in bushfire year \( i \). The population in each state or territory is reported annually in the Australian Historical Population Statistics (from the ABS). The 2008 bushfire year state and territory populations were extrapolated from 2007 values using the average population growth rate over the previous 5 years. Where a bushfire event impacted more than one state or territory, the database provides a geographical breakdown of fatalities so that the data can be normalized separately and added together to determine the \( F_{08} \) numbers. This was similarly the case for the HE data and normalization of them.

d. Validation of normalization methods

Equations (1) and (2) assume that growth in the exposure—number of bushfire-prone dwellings and population in the areas impacted—occurred at the same rate as the growth in total number of dwellings and population for each state or territory. Except for a few particular bushfires, data are not available to allow a more precise estimate of growth in exposed areas over time.

We can get some sense of the relative accuracy of this assumption by comparing state/territory-based dwelling number and population event factors with those derived by weighting equivalent local-level factors by each local area’s proportional contribution to event building damage and fatalities. Urban center/locality (UCL)-based factors were calculated for two of the most damaging historical bushfires: the 1967 Hobart fires and 1983 Ash Wednesday fires. Although UCL-level growth may not necessarily mirror bushfire-prone dwelling and population growth, we expect this to be a more accurate representation than the state/territory-level figures.

The UCL structure is one of the seven interrelated classification structures of the Australian standard geographical classification that groups census collection districts together to form areas defined according to population size (from the ABS). In broad terms, an urban center is a population cluster of 1000 or more people, and a locality comprises a cluster of between 200 and 999 people. The number of dwellings and population in each UCL is reported in census years in the census of population and housing (available from the ABS).

e. ENSO and IOD

There exist multiple definitions for the El Niño and La Niña phases of ENSO, based upon either the Southern Oscillation index (SOI) or various sea surface temperature (SST)-based metrics, but these generally concur for the major El Niño and La Niña events. Here we adopt the Japan Meteorological Agency (JMA) index of 5-month running mean of spatially averaged SST anomalies over a region of the tropical Pacific (4°S–4°N, 150°–90°W). An ENSO year from October through to the following September is then categorized as El Niño (La Niña) if JMA index values are 0.5°C (–0.5°C) or greater (less) for at least 6 consecutive months (including October, November, and December). All other years are classified as neutral. The JMA index for the post-1949 period is based on observed data and, for the years 1925–48, upon reconstructed monthly mean SST fields (Meyers et al. 1999).

In a similar way, IOD events are definition dependent, and we adopt that of Cai et al. (2009a). They define an event using an index of the IOD called the dipole mode index (Saji et al. 1999) in spring (September, October, and November), referenced to the climatological mean over the period 1880–2008. A pIOD (nIOD) event occurs when the index is greater (less) than 0.75 of its long-term standard deviation. Cai et al. (2009a) focused on the spring season as this is when pIOD events peak, and they relate the classification to the following summer season (December, January, and February). All other years are classified as neutral.

The above ENSO and IOD definitions correspond for the worst months for Australian bushfire impacts—December, January, and February—with bushfire years defined earlier as 12-month periods starting 1 July. Classifications according to the above definitions are given in Table 1 (see http://coaps.fsu.edu/jma.shtml). Cai et al. (2009c) noted the recent high frequency of pIOD events, with five occurring during 2002–08.

f. Post–Black Saturday observations

After the Black Saturday fires, Risk Frontiers undertook an aerial reconnaissance for the Kinglake area and Melbourne’s northeastern suburbs, which are interfaced with extensive bushland. On-the-ground surveys were not possible at the time (11 February 2009), with access to many of the impacted areas prohibited while police conducted crime-scene investigations.

Quantitative damage analysis focused on Marysville and Kinglake, the two towns most severely damaged. The main aim was to reveal the spatial pattern of destroyed properties in relation to distance from surrounding bushland boundaries. A Melbourne-based company, Airtech (http://www.airtechaustr.com/), provided 15-cm-resolution, georeferenced postfire imagery captured on 22 and 24 March 2009. These images were manually interpreted, and locations of a total of 1156 destroyed buildings and other surviving structures were digitized. For the distribution
and extent of prefire bushland, we performed various supervised image classifications with the 2.5-m-resolution, orthorectified imagery in the 2009 SPOTMaps series (http://access.spot.com/). It was possible to reliably evaluate the best classification results given the fine resolution of imagery employed and the relative small size of the study area. Once the locations of buildings and bushland boundaries were known, we then calculated distance-based statistics relevant to land use planning and insurance pricing.

3. Results

a. Case studies

We first test the legitimacy of our assumptions that HE and fatalities have increased in proportion to the state/territory-level increase in the total numbers of dwellings and population. Table 2 shows state-based dwelling and population factors (as defined previously) for the 1967 Hobart and 1983 Ash Wednesday bushfires as well as UCL-weighted event factors, calculated by weighting UCL dwelling and population factors by their relative contribution to the total event HE and fatality numbers.

For the Hobart fires, state dwelling and population factors closely mimic their UCL-weighted equivalents (Table 2). Seven UCLs were used to calculate the weighted dwelling factor, with a 60% weight given to the Hobart UCL factor. Only the Hobart UCL was used for the weighted population factor as all 64 fatalities occurred there. The closeness of the state and UCL-weighted factors is not surprising given the size of Hobart as compared with the rest of Tasmania—in 2008 the population in the Hobart UCL stood at approximately one-quarter of the total for the entire state (from the ABS). In other words, the state-based figures are also highly weighted toward Hobart.

In contrast to the Hobart fires, 12 UCLs were used to calculate the weighted dwelling factor for the Ash Wednesday fires, with no one contributing more than 22% of the total building damage. Similar treatment was used for the weighted population factor, where nine UCLs were used with a weighting of not more than 19% applied to each of the contributing UCL factors. All of the UCLs impacted by this event were small relative to the state in which they are located, and the differences are greater than was the case for Hobart, with both state-level factors underestimating growth in dwellings and population at the local level (Table 2).

Table 2 gives some confidence that, while it is possible for state- and local-level normalization factors to diverge, the variation does not appear to be systematic; if anything, it provides some indication that our assumption may be conservative as state-level data may underestimate dwelling growth in exposed areas. It was not possible to derive UCL-weighted factors for the entire analysis period as the UCL structure did not exist prior to the 1966 census (from the ABS).

Table 3 compares the Hobart and Ash Wednesday UCL-weighted normalized HE and fatalities with the data recorded for these events together with those experienced on Black Saturday. The key observation is that the Black Saturday death toll appears to be an aberration: after normalizing the data, the ratio of fatalities to building damage in the Black Saturday fires is more than double that for Hobart and Ash Wednesday. We note that the normalization factors are different for HE and fatalities and so the values shown in the final column in Table 3 are not just a simple arithmetic ratio of the recorded data.

b. Time series of building damage and fatalities

Figures 1a and 2a show time series of the annual aggregated bushfire HE and fatalities recorded in Peril-AUS and the Haynes et al. (2010) database for bushfire years 1925–2008; Figs. 1b and 2b present the corresponding

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**Table 1.** Years 1925–2008 (each beginning 1 Jul) identified as either El Niño or La Niña (ENSO), and pIOD or nIOD (IOD). Other years are classified as neutral for each oscillation.

<table>
<thead>
<tr>
<th>Years</th>
<th>ENSO</th>
<th>La Niña</th>
<th>IOD</th>
</tr>
</thead>
</table>

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**Table 2.** State and UCL-weighted dwelling and population normalization factors for the 2008 bushfire year.

<table>
<thead>
<tr>
<th>Bushfire</th>
<th>State dwelling factor</th>
<th>UCL-weighted dwelling factor</th>
<th>State population factor</th>
<th>UCL-weighted population factor</th>
</tr>
</thead>
<tbody>
<tr>
<td>Hobart 1967</td>
<td>2.0</td>
<td>2.1</td>
<td>1.3</td>
<td>1.1</td>
</tr>
<tr>
<td>Ash Wednesday 1983</td>
<td>1.5</td>
<td>1.8</td>
<td>1.3</td>
<td>1.9</td>
</tr>
</tbody>
</table>
normalized values. Regression analysis on the recorded data reveals increasing trends (Figs. 1a and 2a), more pronounced in the case of Fig. 1a—trends that are reversed in the normalized data (Figs. 1b and 2b). None of these trends are statistically significant at the 10% level, and the overriding impression is of a time series that is dominated by occasional extreme excursions from the mean.

The average annual normalized HE over all years is 301 (Table 4), and the equivalent figure for fatalities is 14. The former is some 3.6 times that determined by McAneney et al. (2009), a difference that arises primarily from normalizing the data. Other factors that influence this difference are 1) the inclusion of other events that had not been previously identified, 2) extending the analysis period to include Black Saturday, and 3) beginning the analysis in 1925 rather than in 1900—the years between 1900 and 1926 for which data exist being characterized by low levels of building damage.

Table 3. Recorded and UCL-weighted normalized HE and fatalities in the 1967 Hobart, 1983 Ash Wednesday, and 2009 Black Saturday fires.

<table>
<thead>
<tr>
<th>Bushfire</th>
<th>HE</th>
<th>Fatalities</th>
<th>Normalized HE</th>
<th>Normalized fatalities</th>
<th>Ratio of normalized fatalities to normalized HE</th>
</tr>
</thead>
<tbody>
<tr>
<td>Hobart 1967</td>
<td>1557</td>
<td>64</td>
<td>3196</td>
<td>70</td>
<td>2.2%</td>
</tr>
<tr>
<td>Ash Wednesday 1983</td>
<td>2253</td>
<td>58*</td>
<td>3958</td>
<td>110</td>
<td>2.8%</td>
</tr>
<tr>
<td>Black Saturday 2009</td>
<td>2852</td>
<td>173</td>
<td>2852</td>
<td>173</td>
<td>6.1%</td>
</tr>
</tbody>
</table>

* Corrected from Haynes et al. (2010).

Figure 1. (a) Annual aggregate HE for bushfire events in PerilAUS for years beginning 1 Jul; (b) as in (a) but with HE normalized to 2008 bushfire year values.
The seemingly anomalous loss of life in the Black Saturday fires and 2008 bushfire year is subject to further scrutiny in Fig. 3, which shows the ratio of annual aggregate normalized fatalities [Eq. (2)] to normalized HE [Eq. (1)] for those bushfire years for which the normalized HE is greater than 600. Adoption of a 600-HE threshold, which conveniently reduces the data to the 10 most damaging years, was done simply to eliminate those years with little or no building damage and/or few or no fatalities. The generally decreasing pattern in Fig. 3 over time is broadly insensitive to the threshold of building damage adopted: a very similar pattern is revealed if a threshold of 100 HE is applied. Of the 10 most damaging years, not since close to the beginning of the analysis period, the 1938 bushfire year, has there been a higher ratio of normalized fatalities to building damage (Fig. 3) than in the 2008 bushfire year. The ratio of total normalized fatalities to HE over the entire analysis period is 4.7%.

### c. ENSO and IOD relationship with normalized bushfire building damage

Table 4 shows the median, average, and standard deviation (of normalized HE) over the 84-yr study period.

<table>
<thead>
<tr>
<th></th>
<th>Median annual normalized HE</th>
<th>Avg annual normalized HE</th>
<th>Std dev of annual normalized HE</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>ENSO</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>El Niño</td>
<td>29 (37)</td>
<td>414 (378)</td>
<td>864 (823)</td>
</tr>
<tr>
<td>Neutral</td>
<td>38 (35)</td>
<td>282 (252)</td>
<td>690 (606)</td>
</tr>
<tr>
<td>La Niña</td>
<td>0 (0)</td>
<td>240 (328)</td>
<td>783 (929)</td>
</tr>
<tr>
<td><strong>IOD</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>pIOD</td>
<td>77</td>
<td>574</td>
<td>1012</td>
</tr>
<tr>
<td>Neutral</td>
<td>10</td>
<td>247</td>
<td>690</td>
</tr>
<tr>
<td>nIOD</td>
<td>0</td>
<td>110</td>
<td>244</td>
</tr>
<tr>
<td>All years</td>
<td>20</td>
<td>301</td>
<td>743</td>
</tr>
</tbody>
</table>
period for ENSO and IOD classified years. There are distinct differences in the median annual and average annual normalized HE for El Niño and La Niña years as there are for pIOD and nIOD years. As expected, the average building damage is highest in El Niño and pIOD years and the median damage in La Niña and nIOD years is zero. The distribution of damage over all years is highly skewed.

The relationship between ENSO and bushfire building damage is reasonably robust, although the strength of the relationship is weakened if an alternative SOI-based definition is applied (Table 4). Under the SOI definition, an El Niño (La Niña) year occurs when the average of June–December monthly SOI values is less (greater) than \(-2.5\) (5) (S. Power 2009, personal communication). The difference between the average building damage in El Niño and La Niña years is substantially reduced under the SOI definition but the median in La Niña years is still zero.

It is important to note the statistics in Tables 4 and 5 are sensitive to building damage in the most destructive of bushfire years. The 10 most damaging years in terms of normalized HE account for almost 80% of total normalized damage, and the ENSO classification of only one of these bushfire years (2008—the fourth largest) differs between the two ENSO definitions: using the SST definition, the 2008 bushfire year is classified as neutral (Table 1) whereas under the SOI definition, it is categorized as La Niña.

Tables 4 and 5 suggest that the IOD is more discriminating than ENSO in relation to normalized bushfire building damage in Australia (the SST definition of ENSO was used in Table 5). The two most damaging combined phases are pIOD/neutral and pIOD/El Niño, which together make up 17 of the 84 bushfire years in the study period (Table 5). The 1938 bushfire year (La Niña year, neutral IOD year) is the only example in which extreme building damage (normalized HE > 1000) occurred in either a La Niña or nIOD year.

As pointed out earlier, bushfire years (beginning 1 July), ENSO years (beginning 1 October), and IOD years (relating to the summer season) do not completely overlap. The effect of this difference is negligible as less than 0.2% of the total normalized HE occurred in the months from 1 July to 30 September inclusive and almost 95% of the total normalized building damage occurred during summer (December, January, February).

d. Post–Black Saturday analysis: Kinglake and Marysville

Destroyed buildings in Kinglake and Marysville were categorized as a function of distance from bushland boundaries, and these data are presented in Fig. 4. A key feature is that about 25% of destroyed buildings were located physically within the bushland boundary, and 60% and 90% were within 10 and 100 m of bushland (Fig. 4). Most buildings in Marysville lay within 200 m of the bushland boundary and, given the wind change that occurred early in the evening on 7 February 2009, would have been subject to ember attack from multiple directions (Victorian Bushfires Royal Commission 2009).

### Table 5. Average annual normalized HE by ENSO (SST definition) and IOD phase for years 1925–2008 (each beginning 1 Jul). Numbers in parentheses are the number of years on which the average is based.

<table>
<thead>
<tr>
<th>ENSO</th>
<th>pIOD</th>
<th>Neutral</th>
<th>nIOD</th>
</tr>
</thead>
<tbody>
<tr>
<td>El Niño</td>
<td>631 (10)</td>
<td>142 (8)</td>
<td>— (0)</td>
</tr>
<tr>
<td>Neutral</td>
<td>645 (7)</td>
<td>239 (35)</td>
<td>110 (6)</td>
</tr>
<tr>
<td>La Niña</td>
<td>39 (2)</td>
<td>358 (10)</td>
<td>110 (6)</td>
</tr>
</tbody>
</table>
4. Discussion

In assuming bushfire-related building damage and fatalities change in proportion to dwelling numbers and population, Eqs. (1) and (2) estimate the number of HE and fatalities in a given event had it occurred under 2008 bushfire year societal conditions. There are other factors not accounted for in the normalization methods, although we expect their influence, particularly on the building damage record, to be minimal relative to societal change. For example, it is likely that some historical bushfires occurred in what were formerly unpopulated areas and thus would have registered no building damage, whereas in these same areas large losses may now be possible. The opposite is also true where original bushlands have been converted to suburbs so that some historical bushfire impacts may now be physically impossible.

Haynes et al. (2010) suggest that a reduction, over time, in the number of people living and working in isolated rural locations would have influenced the fatality data. The effect of this shift was evident in the decreased number of fatalities due to late evacuation, the most common activity at time of death (Haynes et al. 2010). More specifically, Haynes et al. (2010) found a marked decline in those who died while evacuating from working outside and they concluded that this in part explained the absence of a trend in the fatality data (prior to the Black Saturday fires) despite considerable population growth. Notwithstanding these and other qualifications, Figs. 1b and 2b show our best estimates of normalized bushfire building damage and fatalities.

Is the normalized building damage realistic? The average nominal value of a new house (excluding land) in Australia in the 2008 bushfire year was approximately AU$260,000 (from the ABS) so that, in dollar loss terms, the average annual building damage of 301 HE (Table 4) equates to AU$78 million. As noted earlier, this amount excludes building contents, cars, and so on and so will underestimate the property loss, but it does include both insured and uninsured building damage. From an independent dataset, but using a conceptually similar normalization method, Crompton and McAneney (2008) found the average annual insured property loss from weather-related natural disasters between 1967 and 2006 to be around AU$820 million (in 2006 dollars), of which about 12%, or AU$98 million, can be attributed to bushfire. Despite the stated differences, the closeness of these two independent estimates provides some confidence in the method and results. The relationship between normalized building damage and ENSO and the IOD provides additional confidence.

Similar to the result of Cai et al. (2009a) we found normalized Australian bushfire building damage to be more strongly related to the IOD than to ENSO. This is unsurprising as Cai et al. (2009a) follow the Ellis et al. (2004) definition of a significant bushfire, and this incorporates historical impacts (fatalities, property, and livestock losses) rather than meteorological variables or indices. The significant Victorian summer bushfire seasons that the Cai et al. (2009a) study is based upon are therefore correlated with years of high normalized Australian building damage (61% of the total normalized HE occurred in Victoria), at least over the common time period since 1950.

The Black Saturday fires rank fourth in terms of normalized building damage. There were 173 fatalities, more than double the recorded number in any other bushfire event over the analysis period. After normalization, the Black Saturday death toll ranks second to the 1939 Black Friday fires with 214 normalized fatalities. In other words, history suggests that even larger impacts are possible under the climate of past decades. However, this ranking should not detract from the extreme impacts and high ratio of normalized fatalities to building damage in the Black Saturday fires and the need for policy changes to reduce the likelihood of this happening again.

One unequivocal result from our analyses is the absence of any significant trend in normalized HE over time (Fig. 1b). This being the case, a reasonable conclusion at this time, consistent with similar studies summarized by Bouwer (2010), is that it is not possible to detect a greenhouse-gas climatic-change signal in the time series of Australian bushfire building damage once it has been normalized. Such an influence is not ruled out by our analysis, but, if it does exist, it is clearly dwarfed by the magnitude of the societal change and the large year-to-year variation in impacts. Moreover it seems highly implausible that the net effect of other factors such as changes...
in bushfire risk management is being exactly balanced by a greenhouse gas–driven climatic-change influence.

5. Policy implications

This study has shown that increasing building damage due to bushfire in Australia is largely being driven by increasing dwelling numbers and that the impact of greenhouse gas–driven climatic change is not detectable at this time. With this in mind, to reduce the impact of future bushfire events, investments to reduce societal vulnerability need to be made and are likely to bring immediate benefit. Adaptation should be undertaken concurrently with mitigation so that success in addressing bushfire risk in Australia in the short term at least is not misunderstood in terms of obtaining global agreement on reduction of greenhouse-gas emissions.

The Black Saturday tragedy occurred in the face of significant investments (Ashe et al. 2009) and improvements in bushfire risk management and suppression. We take the view that the extreme property losses were in part related to the close proximity of many dwellings to bushland. Chen and McAneney (2004) showed that, although distance to bushland is not the only variable determining bushfire vulnerability, it is demonstrably the most important, with the probability of home destruction decreasing strongly as a function of this distance, a result interpreted as being indicative of ember density and flammability. In the towns of Kinglake and Marysville, where the majority of building damage occurred in the Black Saturday fires, we have shown that 25% of destroyed buildings were literally located within bushland and that 60% were within 10 m of the bushland boundary. Under the extreme conditions prevailing on that day, it is difficult to imagine that homes in the flame zone could have been successfully defended against the combined threats of flames, radiant and convective heating, and embers.

The “prepare, stay, and defend or leave early” bushfire policy, adopted by all Australian fire authorities at the time of the Black Saturday fires, arose on the basis of concerns about the likelihood of large losses of life occasioned by late evacuation (Handmer and Tibbits 2005; Haynes et al. 2010) and the perceived impracticability of evacuating large numbers of people every time severe bushfire conditions exist, circumstances that might arise in some years and some parts of the country for much of the summer. The policy put the actions of residents as central in the protection of lives and property and has the commendable attribute of discouraging an unwarranted dependence upon emergency services. On the other hand, an incorrect interpretation of this policy during Black Saturday may also have contributed to a mistaken belief that homes constructed within or in close proximity to the bushland could be successfully defended against bushfires in extreme conditions.

Our results raise serious questions about land use planning in Australia in relation to bushfire risk. The comparison of state- and UCL-level normalization factors (Table 2) is cause for further concern as it suggests dwelling growth in areas of high bushfire risk may be occurring faster than state averages. We echo the sentiments of McAneney et al. (2009) that, without changes in policy, particularly in land use planning, further bushfire catastrophes are inevitable.

REFERENCES


